



## Group membership, team preferences, and expectations<sup>☆</sup>

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### ABSTRACT

Group membership is a powerful determinant of social behaviour in a variety of experimental games. Its effect may be channelled primarily via the beliefs of group members, or directly change their social preferences. We report an experiment with a prisoner's dilemma with multiple actions, in which we manipulate players' beliefs and show that group identity has a consistent positive effect on cooperation only when there is common knowledge of group affiliation. We also test the robustness of the minimal group effect using three different manipulations: one manipulation fails to induce group identity, and we observe an unsystematic effect of group membership when knowledge of affiliation is asymmetric.

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## 1. Introduction

There is currently a revival of interest among economists in the effect of group membership on individual decision-making.<sup>1</sup> It is well known that people tend to behave more pro-socially when they interact with members of their own group, but become less generous, less trusting, and less cooperative towards individuals who belong to different groups. However, there is less agreement about why this happens, and in which conditions group membership has a significant effect.

The experiment described in this paper extends research on group membership in two directions. (1) Using a two-person public goods game (or multiple-action prisoner's dilemma), it tries to discriminate between two alternative explanations of group identity effects. Does group membership change people's *goals* (by, for example, modifying the argument of their utility function) or does it change people's *expectations* concerning what other individuals will do? (2) The experiment probes the robustness of group effects comparing three different versions of the classic minimal group paradigm (Tajfel et al., 1971).

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<sup>1</sup> See e.g. Akerlof and Kranton (2000, 2010), Sugden (2000), Eckel and Grossman (2005), Bacharach (2006), Cooper and Kagel (2005), Bernhard et al. (2006), Goette et al. (2010), Ruffle and Sosis (2006), Charness et al. (2007), Efferson et al. (2008), Hargreaves Heap and Zizzo (2009), Chen and Li (2009), Sutter (2009), Benjamin et al. (2010) and Zizzo (2011).

While the answer to the first question appears rather straightforward – the effect of group membership is channelled mainly through people's expectations – the results on the second front are less univocal. Minimal group manipulations appear to be fragile, and have unsystematic effects when knowledge of group membership is asymmetric. In some sessions group identity increases transfers to fellow group members, in some it decreases transfers, and in others it has no effect at all.

The paper is organised as follows: Section 2 sketches the theoretical background and briefly reviews the experimental literature. The design of the experiment is illustrated in Section 3, while Section 4 describes and discusses critically the main results. Section 5 concludes with a summary and general comments.

## 2. Literature review

In the classic minimal group experiment Tajfel et al. (1971) divided subjects in two groups using an irrelevant and arbitrary criterion. Subjects then allocated money between random in-group and out-group members, and on average gave more to the former than to the latter. It is noteworthy that subjects sometimes sacrificed resources to increase the difference between in-group and out-group payoffs, that is, they behaved spitefully towards out-group members.<sup>2</sup>

It remains unclear, however, *how* exactly the minimal group design generates higher levels of transfer towards in-group members. Possible explanations can be divided in two broad categories: according to *preference*-based models, group identity transforms the utility functions of individuals who are engaged in a collective task; according to *belief*-based models the group identity manipulation changes their expectations, and via this route modifies behaviour.

The simplest preference-based models introduce other-regarding concerns in the utility function of each individual player. "Social preferences" may be altruistic, egalitarian, reciprocal, spiteful, or may reflect a combination of different motives (Fehr and Fischbacher, 2002; Cooper and Kagel, *in press*). Group identity may change the weight of other-regarding relative to self-interested motives, inducing differential treatment of in-group and out-group members. According to an alternative, less orthodox hypothesis, group identity may cause individuals to focus on the maximisation of a single *team preference* function (Sugden, 2000; Bacharach, 2006). An advantage of this framing effect is that some strategic problems are transformed into parametric decisions, where each individual simply pursues the group's goal by choosing a profile of strategies that maximises collective utility.

Belief-based models in contrast explain the effect of group membership as a manipulation of expectations. In public goods games, for example, individuals with an underlying preference for conditional cooperation must be reassured that others are also willing to contribute. Information about group membership may work as a signal or correlation device that individuals use to coordinate their choices (Bicchieri, 2006; Gintis, 2009). It is crucial however that group affiliation is common knowledge among players. Suppose for example that *i* believes that *j* does not expect her to contribute to the public good. The minimal group paradigm may change *i*'s behaviour by manipulating her beliefs concerning *j*'s expectations. But *i*'s beliefs can change only if *i* learns that *j* knows that *i* knows that they are fellow group members.<sup>3</sup>

A substantial body of evidence confirms the importance of expectations in sustaining pro-social behaviour (Kagel et al., 1996; Haley and Fessler, 2005; Dana et al., 2007; Bicchieri and Xiao, 2009; Ellingsen et al., 2012). None of these experiments, however, focuses specifically on groups. Notable exceptions are Yamagishi and Mifune (2008) and Güth et al. (2009), who have tested the importance of mutual beliefs in dictator's games with group identity. They report significant differences between in-group and out-group allocations *only* when group affiliation is common knowledge.<sup>4</sup> Jin and Yamagishi (1997) similarly studied asymmetric knowledge of group membership in a prisoner's dilemma game. They report higher rates of cooperation only with mutual knowledge of affiliation, but we do not know the details of their design because the original paper was published in Japanese.

In this paper we describe an experiment based on a prisoner's dilemma, that in addition probes different manipulation devices and checks their robustness using post-experimental questionnaires. We manipulate players' *beliefs* and compare conditions with common knowledge of group membership vs. conditions with asymmetric knowledge. In the latter, all players are aware of *their own* group affiliation, but some players do not know the affiliation of the other player (who, in turn, knows that the first player ignores this piece of information). If beliefs are crucial, the difference between in-group and out-group cooperation should be larger in the common knowledge than in the asymmetric knowledge condition. As we shall see, our results confirm the hypothesis that beliefs matter: group membership does not affect cooperation systematically,

<sup>2</sup> Subsequent work in social psychology has explored various alternative methods to induce group identity. See e.g. Tajfel (1982), Brewer and Kramer (1986), Isaac and Walker (1988), Orbell et al. (1988), Dawes et al. (1990) and Kerr and Kauffmann-Gilliland (1994); the social psychology literature is surveyed in Brown (2000) and Hogg and Abrams (2003). We will use the term "minimal group" rather broadly, to include a number of experiments that differ in some respects from Tajfel's. There are various degrees of "minimality", and our experiment probes the robustness of the effect to changes in the manipulation device. Notice also that while Tajfel's subjects engaged in a task that had no payoff consequences for themselves, we follow the experimental economics tradition and study situations where pro-social behaviour has a cost for the decision-maker.

<sup>3</sup> For imagine that only *i* knows about the common group affiliation: since *j* does not know whether she is playing with an in-group or an out-group member, she is unable to infer the correct rule for that situation, and she cannot do better than play randomly. Player *i* as a consequence is also unable to predict the contribution of *j*, and cannot do better than play randomly. The minimal group paradigm should have no significant effect on the average behaviour of experimental subjects in one-shot games with asymmetric information of group membership.

<sup>4</sup> Interestingly, Güth and co-authors report a significant difference between the asymmetric and the common knowledge conditions only when the dictator's beliefs concerning the recipient's expectations are elicited *in advance* of making her decision. When the dictator's attention is not focused on mutual beliefs, in contrast, the asymmetry of information does not seem to matter.

unless both partners are aware of their common affiliation. However, the data also cast doubt on the robustness of the minimal group effect and on subjects' interpretation of the task in the (rather unusual) asymmetric knowledge conditions.

### 3. Experimental design

The basic structure of our experiment is a  $2 \times 3$  design (across subjects) where we manipulate (1) the group membership of two players in a one-shot prisoner's dilemma with multiple actions, and (2) information concerning the group affiliation of the other player. Along the first dimension, we have pairs of subjects belonging to the same group in some sessions (IN-group conditions), and pairs belonging to different groups in other sessions (OUT-group conditions). Along the second dimension, we have sessions with common knowledge of group membership (CK), and sessions with asymmetric knowledge where one player is aware of the group affiliation of both players (AK\_full) while the other one knows her own affiliation but ignores the affiliation of her partner (AK\_partial). In all cases, the prisoner's dilemma game was preceded by a priming session using the minimal group paradigm.

Overall 410 subjects participated in the experiment, drawn from the student population of the University of Trento in the North of Italy. Subjects were recruited using flyers and registered in a dedicated website of the Cognitive and Experimental Economics Laboratory. As they entered the laboratory, subjects were seated randomly at computer desks separated by partitions. An assistant read the instructions aloud and invited subjects to answer six questions to test their comprehension of the experimental task. The assistant then illustrated the correct answer to each question and encouraged further requests of clarification. When all doubts had been dispelled, the experiment began.

*Stage 1* of the experiment consisted of a minimal group manipulation. To cross-check the effect of the group identity manipulation, we used different techniques in different sessions. The manipulations are labelled "Guess", "Bracelets", and "Painters".

**Guess:** subjects were asked to estimate the number of students currently registered at the University of Trento. It was made clear in the instructions that the only purpose of this task was to divide them into separate groups on the basis of a similarity criterion (so that each subject would belong to a group of *similar* individuals).<sup>5</sup> Once they had made their guesses, subjects were told that those whose answers lay above the median would be assigned to the "Yellow" group, and those below the median to the "Red" group. The value of the median was then calculated and each subject's group affiliation was communicated privately.

**Bracelets:** subjects picked randomly a coloured bracelet (Red or Yellow), which they were asked to wear during the experiment. The random draw method has been used by several other experimenters before (following Billig and Tajfel, 1973); the only difference is that we tried to enhance the perception of similarity using a physical marker that remained salient throughout the task.

**Painters:** subjects were asked to evaluate a series of paintings by Vassily Kandinsky, assigning a score from one to ten. The median score was then communicated, and subjects were divided in the Yellow or Red group depending on whether their own score lay above or below the median.<sup>6</sup> To further bolster group identity, subjects engaged in a recognition task (they had to identify the authors of five modern art paintings) which earned them five cents for every correct answer provided by a member of their group (including one's own answers). This device was meant to implement the "common fate" condition that according to social psychologists constitutes an important element of group identity. The results (earnings) of this task were communicated at the end of the experiment.

After group identity had been primed using one of these three techniques, *Stage 2* of the experiment began. Subjects were paired randomly and asked to play a two-person linear public goods game (a prisoner's dilemma with  $n > 2$  discrete actions). Each subject received an endowment of ten euro, to be allocated in units of one euro across two separate accounts. Each unit invested in the "Personal Account" produced exactly one euro for that player. Each unit invested in the "Public Account" was added to those invested by the other player, multiplied by a factor of 1.5, and divided equally between the two players. The production function was therefore

$$\pi = E - c_i + .75 \times (c_i + c_{j \neq i}),$$

where  $E$  is the initial endowment, and  $c_i, c_j$  are the contributions of the two players.

Subjects were asked to allocate the endowment by entering two numbers (one for the Private and one for the Public Account) in separate boxes on their computer screen. As anticipated in the instructions, the screen contained information concerning the affiliation of the other player. In the *common knowledge conditions* (CK) it said "The other player is Yellow [Red]. He/she knows that he/she is Yellow [Red], knows your colour, and knows that you know both players' colours". In the *asymmetric knowledge conditions*, it said either "The other player may be Yellow or Red" (condition AK\_partial), or "The other player is Yellow [Red]. He/she knows his/her colour, but does not know your colour" (condition AK\_full). To make group affiliation salient, we represented it visually using two human-shaped icons coloured in Yellow or Red. When subjects did

<sup>5</sup> The word "similar" was underlined in the instructions, to convey an idea of group homogeneity (according to cognitive psychologists homogeneity is an important factor in group framing). However, the arbitrariness of the mechanism used to divide subjects in groups was totally explicit and no deception was involved.

<sup>6</sup> We used this device to replicate Tajfel's original task as closely as possible, but without deception.

|                           |         | IN  | OUT   |
|---------------------------|---------|---|---|
| Asymmetric Knowledge (AK) | Full    |  |   |
|                           | Partial |  |  |
| Common Knowledge (CK)     |         |  |   |

Fig. 1. Summary of conditions and icons used in the experiment.

not know their partner's affiliation, the icon on the right-hand side was coloured in grey and carried a large question mark. When subjects were told that their affiliation had not been disclosed to their partner, the icon on the right-hand side was appropriately coloured (Red or Yellow) but carried a balloon with a question mark to signify the partner's lack of information. Fig. 1 summarises the various treatments and displays some of the icons that we used.

After all subjects had made their decision, they answered a short questionnaire that elicited their subjective experiences of participation in the experiment. One question in particular probed their feeling of identification with the group, and will be discussed in more detail in the next section. At this point each participant received feedback about the money she had earned, filled in a brief questionnaire requesting generic information about age, gender, university degree, etc., and was paid privately in cash (the average earning was roughly 12 euros).

#### 4. Results

Table 1 includes the relative frequency of contribution choices across the experimental conditions. On the rows we report the results of the three beliefs conditions (CK, AK.partial, AK.full), while on the columns we report some statistical indicators, organised according to the two treatments (IN and OUT). We also disaggregate the data according to the manipulation device that we have used (Guess, Bracelets, Painters), and then aggregate all the data in a pooled sample (at the bottom).

We start with a broad description of the main patterns, and move subsequently to more detailed statistical analysis. It is immediately clear looking at the pooled sample that with common knowledge of group affiliation (CK) there is a difference of behaviour across the IN and OUT conditions: in-group pairs contribute more. In contrast, IN/OUT contributions barely differ

Table 1  
Contribution levels across conditions and treatments.

|            | IN |       |       |         | OUT |       |       |         |
|------------|----|-------|-------|---------|-----|-------|-------|---------|
|            | N  | Avg   | Med   | Std dev | N   | Avg   | Med   | Std dev |
| Guess      |    |       |       |         |     |       |       |         |
| AK.partial | 32 | 4.438 | 4.500 | 2.873   | 31  | 3.645 | 2.000 | 3.498   |
| AK.full    | 32 | 5.906 | 5.000 | 3.315   | 31  | 3.419 | 3.000 | 3.233   |
| CK         | 32 | 5.188 | 5.000 | 3.031   | 32  | 3.812 | 3.000 | 2.934   |
| Bracelets  |    |       |       |         |     |       |       |         |
| AK.partial | 16 | 4.500 | 3.500 | 3.812   | 18  | 5.222 | 5.000 | 2.981   |
| AK.full    | 16 | 5.250 | 5.000 | 3.022   | 18  | 5.333 | 5.000 | 3.531   |
| CK         | 20 | 5.100 | 4.000 | 3.194   | 20  | 4.500 | 4.000 | 2.947   |
| Painters   |    |       |       |         |     |       |       |         |
| AK.partial | 20 | 4.050 | 4.500 | 3.332   | 18  | 4.056 | 4.000 | 2.775   |
| AK.full    | 20 | 3.550 | 4.000 | 2.685   | 18  | 7.167 | 8.500 | 3.185   |
| CK         | 20 | 6.650 | 6.500 | 3.407   | 16  | 4.000 | 3.000 | 4.000   |
| Pooled     |    |       |       |         |     |       |       |         |
| AK.partial | 68 | 4.338 | 4.000 | 3.203   | 67  | 4.179 | 4.000 | 3.205   |
| AK.full    | 68 | 5.059 | 5.000 | 3.195   | 67  | 4.940 | 5.000 | 3.613   |
| CK         | 72 | 5.569 | 5.000 | 3.210   | 68  | 4.059 | 4.000 | 3.181   |

when knowledge is asymmetric. While this is to be expected for those players who know their own affiliation but do not know the affiliation of their partner (AK\_partial), the results of fully informed players (AK\_full) are theoretically interesting: the data suggest that knowing that the other player belongs to your own group is insufficient, by itself, to induce higher levels of cooperation. This is *prima facie* evidence that the minimal group manipulation influences the *expectations* of players, rather than changing their preferences directly.

If we look at the three manipulations separately (Guess, Bracelets, and Painters), we notice that the effect of group membership is strikingly consistent in the CK conditions. In contrast, there is more variation in the AK conditions. These differences must be taken with a pinch of salt given the limited number of observations, but nevertheless offer insights in the decision process and invite some methodological reflections that will be outlined later. Consider that in the Guess manipulation fully informed players (AK\_full) contribute more when they are matched with an in-group than with an out-group player. But in the Painters manipulation the opposite is true: AK\_full subjects contribute more when they are matched with an out-group than when they are matched with an in-group player. Data under the Bracelets manipulation fall roughly in between: IN/OUT behaviour is practically indistinguishable.

To confirm these first impressions, we run a series of Wilcoxon Rank Sum tests across all conditions.<sup>7</sup> The only significant difference in the pooled data concerns the CK.IN and CK.OUT conditions ( $p = .006$ ). The discrimination seems to originate in a higher propensity to cooperate with in-group members, rather than in a tendency to free ride more with out-group members. This is apparent if we compare common knowledge conditions with asymmetric knowledge conditions: CK.IN is significantly different from AK\_partial.IN ( $p = .028$ ) and AK\_partial.OUT ( $p = .008$ ), while CK.OUT does not differ from AK\_partial.IN ( $p = .594$ ) and AK\_partial.OUT ( $p = .870$ ).<sup>8</sup> It is noteworthy that while the CK.IN vs. CK.OUT difference is significant or close to significance in the Guess and Painters conditions ( $p = .069$  and  $p = .038$ , respectively), it fails to reach significance in the Bracelets condition ( $p = .621$ ). However, this is the only condition where no difference whatsoever is observed across all comparisons, which suggests that the manipulation device failed to generate group identity. We will double-check this hypothesis shortly using some questionnaire data.

The other anomalous results concern the AK\_full conditions. While in Bracelets fully informed subjects do not discriminate significantly between IN and OUT partners ( $p = 1.000$ ), in Guess and in Painters they do, but in opposite ways: in Guess they cooperate more when matched with in-group partners ( $p = .005$ ), while in Painters AK\_full subjects cooperate more with out-group players ( $p = .002$ ). In Painters, however, the anomalously high values observed in AK\_full.OUT are due to a single experimental session where ten subjects contributed on average 2.8 tokens to in-group partners and 9.1 tokens to out-group partners. In the other session that we ran with the same manipulation device, the average contributions were 4.3 and 4.75 respectively, very much in line with what we observed in the experiment overall. We conjecture that if we had had the chance to run other sessions, the anomaly of the Painters manipulation would have disappeared.<sup>9</sup>

To improve our understanding of the determinants of contribution we also run a regression estimation.<sup>10</sup> Contribution to the Public Account is taken as the dependent variable, and alternative experimental conditions are considered as explanatory factors. More precisely, CK is set equal to one when a participant is in the common knowledge condition and zero otherwise; AK\_full is equal to one when a participant is in the asymmetric knowledge condition and zero otherwise. The interaction between these two explanatory factors and the group membership of the other player is also considered. The variable IN is set equal to one when the other player has the same colour as the decision maker and equal to zero when the colours are different.<sup>11</sup> We also add a few control variables to the regression: age captures the age of the decision maker; female captures the gender of the decision maker; freshman is equal to one when the decision is made by a first-year student and equal to zero otherwise; economics is equal to one when the student is an economics major and equal to zero otherwise.

The regression outcomes of Table 2 show that when data from the three experiments are pooled together (column [1]), subjects in the CK condition tend to cooperate more with IN than OUT members ( $CK \times IN$ ). Contributions in CK.OUT do not differ from those in AK\_partial, which we take as our baseline condition. However, LinHyp.1 tells us that contributions to the public good are higher in CK.IN than in the baseline (AK\_partial). Concerning AK\_full, no significant difference is observed between in- and out-group conditions. Furthermore, only a marginally significant difference is registered between contributions in AK\_full.IN and contributions in the baseline (AK\_partial).

If we analyse each manipulation separately, heterogeneity of behaviour emerges once again. While contributions in condition CK are never lower when interacting in the in-group condition than when interacting in the out-group condition, in conditions with asymmetric knowledge (AK) behaviour is more volatile. Under the Guess manipulation, a highly significant positive difference between in- and out-group contributions is observed. Moreover, the contributions in condition IN are

<sup>7</sup> Since there are more than sixty tests, we report only the most relevant results here. A complete table with all  $p$ -values can be found in the working paper version of this article (Guala et al., 2012).

<sup>8</sup> Notice that in theory the data from AK\_partial.IN and AK\_partial.OUT conditions could be pooled together, given that subjects received exactly the same information. The statistical tests confirm that there are no significant behavioural differences.

<sup>9</sup> This would have made recruiting very challenging because of the large number of participants that had already taken part in previous sessions of the experiment.

<sup>10</sup> We used a Tobit regression analysis to account for the censoring at 0 and 10 in the dependent variable. Since the data are discrete, we compute robust standard errors.

<sup>11</sup> In the reported estimation the variable IN is present only in interaction with dummies CK and AK\_full. This specification allows us to obtain the condition AK\_partial as a baseline, with no distinction between in- and out-group conditions.

**Table 2**

Contribution to public good (Tobit regression).

|                | Coeff. (robust standard error) |                             |                             |                              |
|----------------|--------------------------------|-----------------------------|-----------------------------|------------------------------|
|                | [1] Pooled                     | [2] Guess                   | [3] Bracelets               | [4] Painters                 |
| Intercept      | 2.552 (1.010) <sup>*</sup>     | 3.776 (2.940)               | 4.646 (1.427) <sup>**</sup> | −0.847 (5.262)               |
| CK             | −0.469 (0.670)                 | −0.500 (0.897)              | −0.973 (0.163)              | −0.297 (1.563)               |
| AK.full        | 0.820 (0.736)                  | −1.013 (1.006)              | −0.224 (1.229)              | 4.407 (1.462) <sup>**</sup>  |
| CK × IN        | 2.192 (0.779) <sup>**</sup>    | 1.661 (0.990) <sup>°</sup>  | 0.422 (1.415)               | 4.039 (1.899) <sup>*</sup>   |
| AK.full × IN   | 0.483 (0.835)                  | 3.691 (1.214) <sup>**</sup> | 0.823 (1.449)               | −4.570 (1.507) <sup>**</sup> |
| Age            | 0.076 (0.039) <sup>°</sup>     | 0.007 (0.122)               | 0.011 (0.030)               | 0.243 (0.232)                |
| Female         | −0.609 (0.468)                 | −0.211 (0.649)              | −1.981 (0.866) <sup>*</sup> | −1.809 (0.956) <sup>°</sup>  |
| Freshman       | 0.353 (0.518)                  | −0.422 (0.762)              | 0.956 (0.988)               | 2.111 (1.253) <sup>°</sup>   |
| Economics      | 0.086 (0.487)                  | 0.285                       | 2.723 (1.044) <sup>*</sup>  | −0.954 (0.924)               |
| LinHyp.1       | $F(1,400) = 6.52^*$            | $F(1,181) = 1.67$           | $F(1,99) = 0.19$            | $F(1,104) = 6.58^*$          |
| LinHyp.2       | $F(1,400) = 3.74^°$            | $F(1,181) = 7.00^{**}$      | $F(1,99) = 0.23$            | $F(1,104) = 0.03$            |
| <i>N</i>       | 408                            | 189                         | 107                         | 112                          |
| Left censored  | 63                             | 29                          | 13                          | 21                           |
| Right censored | 60                             | 25                          | 16                          | 19                           |
| <i>F</i>       | 2.15 <sup>*</sup>              | 1.84 <sup>°</sup>           | 2.14 <sup>*</sup>           | 3.35 <sup>*</sup>            |

LinHyp.1: CK + CK × IN = 0.

LinHyp.2: AK.full + AK.full × IN = 0.

<sup>\*</sup> Significance level: 0.05.<sup>\*\*</sup> Significance level: 0.01.<sup>°</sup> Significance level: 0.1.

higher than those in the baseline. In Bracelets the treatments do not produce any significant effect. In Painters we have the same surprising pattern discussed above: contributions in condition AK.full.OUT are significantly higher than in condition AK.full.IN and in the baseline.

To put these data in perspective, we discuss briefly the results of a short questionnaire that subjects completed *after* they had chosen their contribution in the Bracelets and Painters manipulation conditions.<sup>12</sup> The questionnaire included several questions concerning the subjective experience of participating in the experiment. The first one, crucially, asked: “During the interaction did you feel that the two participants were like one group or like separate individuals?” (1 = group, 2 = individuals). In the Painters manipulation, 60 percent of participants answered positively (they felt as a group) in the CK.IN condition, compared to only 12.5 percent in CK.OUT. This difference is highly significant ( $p = 0.006$ , Fischer Exact Test). The answers in AK.full.IN and AK.full.OUT were practically indistinguishable, in contrast (25 vs. 27.8 percent,  $p = 1.000$ , FET), which again confirms the importance of common knowledge of affiliation for the creation of group identity. The picture in Bracelets is more blurred, which explains why the behavioural patterns are also rather uninteresting under this treatment. In CK.IN, 40 percent of subjects felt as group, compared to 35 percent in CK.OUT ( $p = 1.000$ , FET); in AK.full the frequencies were 50 and 44.4 percent respectively ( $p = 1.000$ , FET).

Questionnaire data support two important methodological points: first of all, the manipulation treatments that we administered had variable effects. We can say with a high degree of confidence that the Painters and, probably, the Guess devices induced group identity effectively, while the Bracelets manipulation did not. This failure may be attributed to the blatantly artificial process of group formation (random draw), and to the fact that subjects in this condition did not engage in any common task before playing the public goods game. The data should make one pause and reflect on all those experiments that have used the “lightest” version of the minimal group design. It is possible that the results reported in the literature suffer from publication bias, and it would be desirable to publish failed replications in the future (see Maniadis et al., 2011).

The second point concerns the anomalous patterns of behaviour observed in the Painters treatment. Recall that subjects with full information in the AK condition cooperated more with out-group than with in-group subjects. The anomalous data are concentrated in one specific session: if we disaggregate the questionnaire results, we find that indeed only 10 percent of subjects in that session identified with in-group members, compared with a strange 30 percent identification with out-group members. In the other session, the data are much more in line with the rest of the experiment (40 percent identification with in-group members vs. 25 percent with out-group members). We should therefore conclude that the manipulation had a bizarre effect in a single session, which should be discounted from the overall results observed in the experiment.

Are these mere statistical anomalies due to low numerosity, or is there something more to be said? It is not surprising, in our view, that we observed the greatest variance of results in the AK conditions. Contrary to the relatively straightforward situation faced in the CK sessions, in AK the experimental subjects receive contradictory messages: first they are given information about group membership. This information is highly salient, and is probably interpreted as a cue for the decisions they will make in the public goods game. At the same time, however, the signal in AK is not common knowledge and cannot

<sup>12</sup> We do not have questionnaire data about the Guess manipulation, unfortunately, because the idea of a manipulation check was suggested by a referee after these sessions had already been done.

be used as a correlation device (Bicchieri, 2006; Gintis, 2009). Subjects probably perceive the contradictory nature of the signal, and struggle to make sense of it in an unusual experimental environment.

This raises important questions concerning the very nature of the minimal group effect. Right from the beginning, Tajfel's paradigm was criticised for its "artificiality". In particular, some critics argued that subjects reacted to an obvious experimental demand to modulate cooperation according to group membership. Responding to his critics, Tajfel agreed with the premise of this argument, but pointed out that not all demand effects are artificial (in the sense of lacking a counterpart outside of the laboratory):

what was no more than a hint from the experimenters about the notion of 'groups' being relevant to the subjects' behaviour had been sufficient to determine, powerfully and consistently, a particular form of intergroup behaviour. [...] The problem then must be restated once again in terms of the need to specify why a certain *kind* of intergroup behaviour can be elicited so much more easily than other kinds [...]. [T]he subjects structured the situation for themselves as one involving relations between groups, and [...] they behaved in ways similar to those habitual to them in situations of this kind. (Tajfel, 1982: pp. 235–236)

Tajfel's considerations may be extended to all experiments that study social norms in the laboratory. To observe the effects of social norms, an experimenter must create expectations about conformity to a behavioural rule that is deemed appropriate to the situation. All framing effects (like those investigated by Eckel and Grossman (2005), for example, or Ellingsen et al. (2012)) exploit a demand effect in this sense. The scientific interest of these studies lies in the hypothesis – which is a priori plausible and may be confirmed by field data – that the experimental manipulation cues behavioural rules that have been "imported" in the laboratory from the outside world. If this is the case, the experimental results have external validity *because* the subjects have complied with the demand (rather than in spite of it).<sup>13</sup>

Our results add another important nuance to this methodological point: the key mechanism linking group identity with cooperation is subjects' concern about the expectations of *their peers*. Our data, in other words, indicate that subjects are not particularly concerned about the expectations of the experimenter, since removing mutual knowledge of group membership tends to make the effect disappear. The minimal group should be considered a "peer demand effect", rather than an "experimenter's demand effect", which makes one think that the validity of group identity extends beyond laboratory walls.

## 5. Summary and conclusion

When they choose their contribution level to a public good, subjects tend to contribute more if they are matched with an in-group member than if they are matched with an out-group member. Thus, pro-social cooperative behaviour is affected by group affiliation. Systematic discrimination however takes place only if both subjects have access to information about the group membership of the other player. When knowledge is asymmetric, fully informed participants do not cooperate more with in-group than with out-group subjects. Overall their average level of contribution is statistically indistinguishable from that of partially informed subjects. The message is that the minimal group paradigm acts primarily on individual beliefs, and through this channel modifies behaviour in games of cooperation. All theories that postulate a direct link between group identity and preferences – unmediated by mutual beliefs – are therefore refuted by the evidence.

The results of this experiment however raise important questions regarding the robustness of minimal group effects. Out of three attempted manipulations, one certainly failed to induce significant discrimination and must be taken as a warning against attempts to ground group identity on mere random labelling. The other two manipulations successfully replicated the classic in–out effect with common knowledge of group membership, but produced unsystematic effects in the asymmetric conditions. This variance is probably due to subjects' attempts to apply in an unusual context a behavioural norm that is appropriate for situations where group affiliation is common knowledge among players.

## Appendix A. Supplementary data

Supplementary data associated with this article can be found, in the online version, at <http://dx.doi.org/10.1016/j.jebo.2012.12.003>.

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<sup>13</sup> For general discussions of the problem of external validity in experimental economics see e.g. Guala (2005), Schram (2005), Levitt and List (2007) and Bardsley et al. (2009).

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